

Robustness of OLS to sample removals: theoretical analysis and implications

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Joint work with
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- Revisit the real data example
- Insights and implications

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Assume:

p = number of features $< n$ = number of samples

X - $n \times p$ matrix of all sample is of full rank p

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where

$$\hat{\beta} = (X^\top X)^{-1} X^\top \mathbf{y}$$

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In the statistics literature: long history of works under umbrella of
Robustness Diagnostics

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In past few years, several researchers noted that in various datasets, removing a *small* number $k \ll n$ of (specifically chosen) training samples leads to *large* changes in the learned model,

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Example: For OLS, let $\mathcal{S} \subset [n]$ be remaining set of samples after removal of k specific samples, $|\mathcal{S}| = n - k$,

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If for a specific coefficient $j \in [p]$, with $k \ll n$ samples removed,

$$\hat{\beta}_j > 0 \quad \text{but} \quad (\hat{\beta}_{\mathcal{S}})_j < 0$$

our trust in the model may be questionable

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Broderick et al, 2020, considered a more stringent (worst-case) form of robustness, called **robustness auditing**:

estimated parameters / predictions need to be robust to removal of *any* subset of k samples

Most influential subset selection problem

[Rubinstein and Hopkins 25']

Definition (OLS): Robustness of $\hat{\beta}$ to k sample removals in a fixed direction $\mathbf{v}$ is measured by

$$\Delta_k(\mathbf{v}) = \max_{\mathcal{S} \subseteq [n], |\mathcal{S}|=n-k} \langle \hat{\beta} - \hat{\beta}_{\mathcal{S}}, \mathbf{v} \rangle$$

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Key Questions:

- Practical: For a given dataset and a given k , how large can $\Delta_k(\mathbf{v})$ be ?
- Theoretical: How large should we expect $\Delta_k(\mathbf{v})$ to be under reasonable assumptions about the data?

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Several of these works show that on various datasets, OLS is *not* robust to removal of even just a handful of samples.

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$x_2, \dots, x_{17}$ additional covariates of each household.

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$x_2, \dots, x_{17}$ additional covariates of each household.

In non-poor households, β_1 captures (indirect) effect of the Progresas aid program.

Cash Transfers Dataset

Period	Poor	n	$\hat{\beta}_1$	AMIP
8	Y	10781	16.53	225
8	N	4543	-5.53	5
9	Y	9489	28.65	321
9	N	3769	23.19	21
10	Y	10368	32.52	570
10	N	4191	21.12	26

6 rows: three time periods $\times$ two groups (poor/non poor)

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Column 5: size of smallest subset found by AMIP
whose removal changes the sign of $\hat{\beta}_1$

Some Questions

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- Implications for practitioners

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- Model 1: General P with mild regularity conditions

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Analyze two models for $P(\mathbf{x}, y)$:

- Model 1: General P with mild regularity conditions
- Model 2: Gaussian Linear Model.

Model 1 - General

$P(\mathbf{x}, y)$ - probability distribution over $\mathbb{R}^{p+1}$ such that

1. $\mathbf{x}$ is sub-Gaussian, zero mean, and has positive-definite covariance $\Sigma = \mathbb{E}[\mathbf{x}\mathbf{x}^\top]$
2. y is sub-Gaussian

Model 2 - Gaussian Linear

$y = \beta^\top \cdot \mathbf{x} + \varepsilon$, where $\beta \in \mathbb{R}^p$ is deterministic and

1. $\mathbf{x} \sim \mathcal{N}(0, \Sigma)$, $\Sigma \succ 0$ is positive definite
2. ε is sub-Gaussian, mean zero, and independent of $\mathbf{x}$

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Under Gaussian-Linear Model 2, $\beta^{\text{OLS}} = \beta$

Sub-Gaussian norm of y :

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Sub-Gaussian norm of vector $\mathbf{x}$:

$$\|\mathbf{x}\|_{\psi_2} = \sup_{\mathbf{v} \in \mathbb{S}^{p-1}} \|\mathbf{v}^\top \mathbf{x}\|_{\psi_2},$$

Non-asymptotic bound

$(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)$ i.i.d. from Model 1.

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Theorem

Assume $k \leq n/2$, $p \leq n - k$, and define

$$\eta = \|\Sigma^{-1/2}\| \left(1 + \|\Sigma^{-1/2}\mathbf{x}\|_{\psi_2}^4\right) \|y\|_{\psi_2}^2.$$

There exist constants $C, c > 0$ such that with probability at least $1 - 7(n/k)^{-ck}$,

$$\max_{S \subseteq [n]: |S| \geq n-k} \|\hat{\beta}_S - \hat{\beta}\| \leq C\eta \sqrt{\frac{k \log(en/k)}{n-k}}$$

for all p, n, k satisfying

$$C \|\Sigma^{-1/2}\mathbf{x}\|_{\psi_2}^2 \sqrt{\frac{p}{n-k}} \leq 1.$$

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Theorem

$(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)$ i.i.d. from Model 1.

Assume that as $n \rightarrow \infty$, $\limsup \|\Sigma^{-1/2}\mathbf{x}\|_{\psi_2}^2 \sqrt{p/n} < c$, where $c > 0$ is an absolute constant. Then, if $k/n \rightarrow 0$,

$$\max_{S \subseteq [n], |S| \geq n-k} \|\hat{\beta}_S - \hat{\beta}\| \xrightarrow{p} 0$$

OLS Robustness Under Model 1

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In fact, theorem allows $p \rightarrow \infty$ provided that $p/n \rightarrow 0$, as in this case the condition of the theorem is satisfied.

Robustness measure $\Delta_k(\mathbf{v})$ converges to zero, uniformly in $\mathbf{v}$:

$$\Delta_k(\mathbf{v}) = \max_{S \subseteq [n], |S|=n-k} \langle \hat{\beta} - \hat{\beta}_S, \mathbf{v} \rangle \leq \max_{S \subseteq [n], |S| \geq n-k} \|\hat{\beta}_S - \hat{\beta}\|.$$

Hence, if $k/n \rightarrow 0$,

$$\sup_{\mathbf{v} \in \mathbb{S}^{p-1}} \Delta_k(\mathbf{v}) \xrightarrow{p} 0.$$

Theorem

Under same conditions above, and additional assumption that $\kappa(\Sigma)$ and $\|\beta\|$ remain bounded, as $n \rightarrow \infty$ and $(p + k)/n \rightarrow 0$,

$$\max_{\mathcal{S} \subseteq [n], |\mathcal{S}| \geq n-k} \|\hat{\beta}_{\mathcal{S}} - \beta\| \xrightarrow{p} 0.$$

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Maximum number of samples that can be removed while keeping OLS error rate:

$$k \log \left(\frac{n}{k} \right) \ll \sqrt{p}$$

$(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)$ i.i.d. from Gaussian Linear Model 2.

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Theorem

There exist constants $C, c > 0$ such that if

$$Ck \leq n, \quad t \geq 0, \quad \text{and} \quad \sqrt{\frac{p}{n-k}} + t < c,$$

then with probability at least $1 - 4e^{-c(n-k)t^2}$,

$$\max_{\substack{S \subseteq [n], \\ |S| \geq n-k}} \|\hat{\beta}_S - \beta\| \leq C \|\Sigma^{-1/2}\| \|\varepsilon\|_{\psi_2} \left(\frac{k \log n}{n-k} + \sqrt{\frac{p}{n-k}} + t \right)$$

Robustness under Linear Model

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Under linear model, OLS can tolerate the removal of significantly more samples than under the general model

On Theoretical Guarantees for ACRE

[Rubinstein and Hopkins, ICLR, 25']

ACRE= Algorithm for Certifying Robustness Efficiently

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For a fixed $\mathbf{v}$, ACRE computes upper and lower bounds $U_k(\mathbf{v})$ and $L_k(\mathbf{v})$ such that without any modeling assumptions,

$$L_k(\mathbf{v}) \leq \Delta_k(\mathbf{v}) \leq U_k(\mathbf{v}).$$

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$$K = \tilde{\Theta}\left(\min\left(\frac{n}{\sqrt{p}}, \frac{n^2}{p^2}\right)\right)$$

such that for all $k \leq K$, with high probability,

$$\frac{U_k(\mathbf{v})}{L_k(\mathbf{v})} = 1 + \tilde{O}\left(\frac{p + k\sqrt{p}}{n}\right).$$

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$$\frac{U_k(\mathbf{v})}{L_k(\mathbf{v})} = 1 + \tilde{O}\left(\frac{p + k\sqrt{p}}{n}\right).$$

When $p + k\sqrt{p} \ll n$, the upper and lower bounds $U_k(\mathbf{v})$ and $L_k(\mathbf{v})$ are tight, so ACRE accurately measures robustness to removals.

On Theoretical Guarantees for ACRE

Comparison to our theoretical results:

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Under model 2, OLS is *provably* robust to removals in a *broader parameter regime* $p + k \ll n$.

On Theoretical Guarantees for ACRE

Comparison to our theoretical results:

Under model 2, OLS is *provably* robust to removals in a *broader parameter regime* $p + k \ll n$.

Open Question: whether the upper and lower bounds of ACRE remain tight in more general regimes, in particular where OLS is non-robust.

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Proof Techniques

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Open Question:

Derive robustness guarantees for other models

Non-robustness for $k \propto n$

$(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)$ i.i.d. from Model 2. Set $\alpha = k/n$.

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Theorem

Fix $\mathbf{v} \in \mathbb{S}^{p-1}$. Assume that $p < k$ and $\gamma = p/(n - k) < 1/4$.

There exist absolute constants $C, c > 0$ such that if $\alpha = k/n \leq c$, then with probability at least $1 - 17e^{-c(n-k)t^2}$

$$\Delta_k(\mathbf{v}) \geq \|\Sigma^{-1/2} \mathbf{v}\| \left(\mathbb{E}[\varepsilon z \mathbb{1}(\varepsilon z > q_{1-\alpha+t})] - \frac{C(t + \sqrt{\gamma})}{(C - \sqrt{\gamma} - t)^2} \right),$$

for any $t \in (0, \min\{\alpha, 1/2 - \alpha\})$. Here, $z \sim \mathcal{N}(0, 1)$ is independent of ε , and $q_{1-\alpha+t}$ is the $(1 - \alpha + t)$ -quantile of εz .

Robustness and Consistency Regimes for OLS / Model 2

Region	$k, p \ll n$	$k \ll n, p \asymp n$	$k \asymp n, p \ll n$	$k \asymp p \asymp n$
Robust	✓	✓	×	×
Consistent	✓	×	✓	×

Cash Transfers Dataset Revisited

Response Y = total household consumption in pesos

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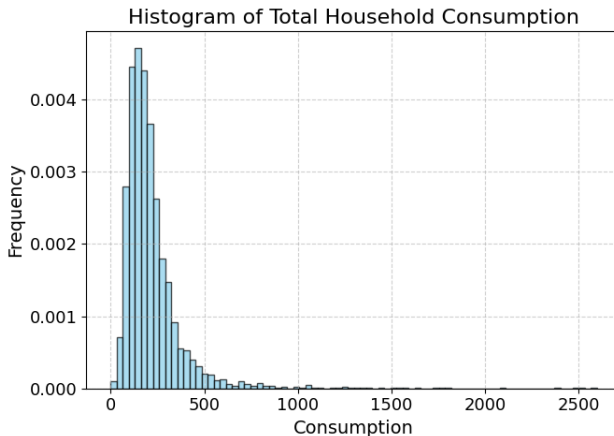
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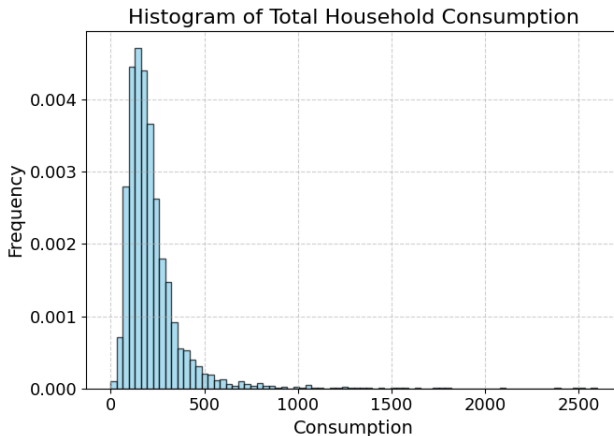
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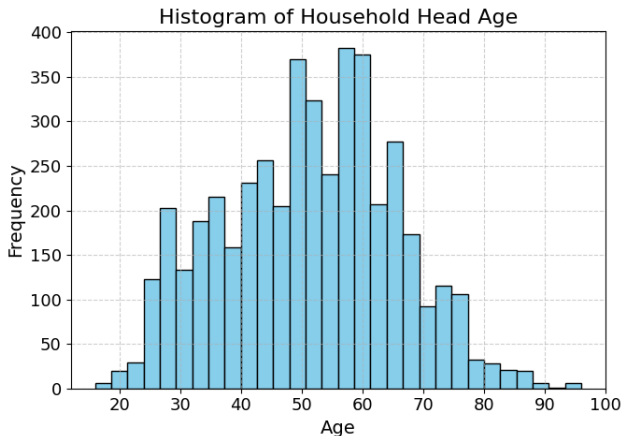
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8	Y	10781	16.53	225	170	126	48	12
8	N	4543	-5.53	5	219	172	29	5
9	Y	9489	28.65	321	176	182	48	15
9	N	3769	23.19	21	226	273	20	9
10	Y	10368	32.52	570	172	156	56	13
10	N	4191	21.12	26	217	267	19	7

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Y is extremely heavy tailed

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Perhaps not surprising few samples suffice to reverse sign $\hat{\beta}_1$

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$\tau > 0$ controls the transition from squared loss to absolute loss.
In our experiments we took $\tau = 1$.

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Assuming AMIP approximation is accurate
Huber regression substantially more robust

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Thank You !